

Lecture 5 | 31.03.2026

Linear regression model with (linear) interactions

Overview: Multiple regression model

- Mathematical relationship between a continuous dependent variable Y and a set of explanatory (independent) variables $\mathbf{X} = (X_1, \dots, X_p)^\top$ (may be continuous, binary, or categorical – or an arbitrary combination)
- Typically expressed for some general function $f : \mathbb{R}^p \rightarrow \mathbb{R}$ but for the linear regression model we use more specific notation of the form

$$\begin{aligned}
 Y &= \beta_0 + \beta_1 X_1 + \beta_{p-1} X_{p-1} + \varepsilon & = \mathbf{X}^\top \boldsymbol{\beta} + \varepsilon \\
 E[Y|\mathbf{X}] &= \beta_0 + \beta_1 X_1 + \beta_{p-1} X_{p-1} & = \mathbf{X}^\top \boldsymbol{\beta}
 \end{aligned}$$

where $f : \mathbb{R}^p \rightarrow \mathbb{R}$ is linear in the unknown parameters $\beta_1, \dots, \beta_p \in \mathbb{R}$

- The corresponding model for a random sample $\{(Y_i, \mathbf{X}_i); i = 1, \dots, n\}$ drawn from some joint distribution function $F_{(Y, \mathbf{X})}$ takes the form

$$Y_i = \mathbf{X}_i^\top \boldsymbol{\beta} + \varepsilon_i, \text{ for } i = 1, \dots, n, \text{ or, alternatively } \mathbf{Y} = \mathbb{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$$

where we assume (by default) the presence of the intercept parameter $\beta_0 \in \mathbb{R}$ in the model (i.e., $X_{i0} = 1$ almost surely)

Quantifying the effect of X on Y

- ❑ One of the main goals of the regression model (regression analysis in general) is to quantify the effect of some given explanatory variable on Y
- ❑ Formally, the explanatory variables in $\mathbf{X} \in \mathbb{R}^p$ typically have an effect on the whole (conditional) distribution of Y ... however, for simplicity, we only focus on some distributional characteristics of this distribution...
- ❑ A typical characteristic related to the linear regression model is the conditional mean/expectation of Y given \mathbf{X} . Therefore, the effect of \mathbf{X} on Y is also explained/interpreted in terms of the corresponding difference of the conditional expected values for two different values of \mathbf{X}
- ❑ The evaluation of the effect may be quantitative (in terms of the estimation of an unknown parameter) or it can be qualitative (in terms of a formal statement whether the effect is statistically important or not)

Interpretation: Association vs. causality

↔ the regression model is typically a model that explains only an association (relationship) between two (or more) subpopulations that differ with respect to the values of the explanatory covariates

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❑ Associative interpretation

- ❑ Comparing two sub-populations that differ wrt to different values of X
- ❑ Interpreting the effect of X in terms of the comparison of two subjects

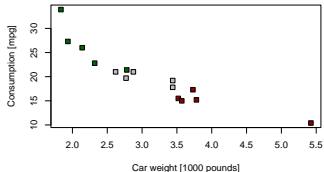
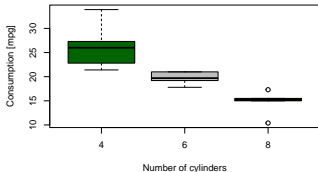
❑ Causal interpretation

- ❑ Comparing the same sub-population before and after the change in X
- ❑ Interpreting the effect of X in terms of a change within the subject

↔ it is a common mistake that the associative regression model is (unintentionally) interpreted as a causal model... however, for a causal interpretation we usually need much stricter assumptions (a randomized trial for instance)

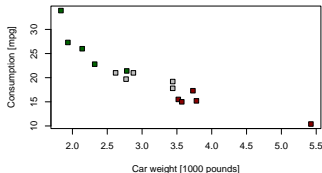
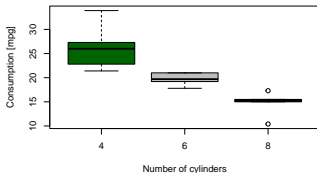
Example: Simple regression model variants

- Consumption of 15 US cars (given the number of cylinders and the car's weight)
(5 cars with 4 cylinders, 5 cars with 6 cylinders and 5 cars with 8 cylinders)



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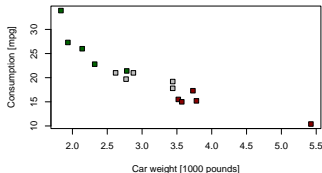
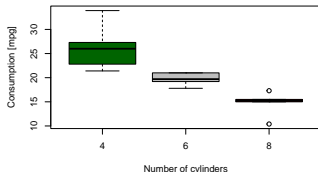


```
lm(formula = mpg ~ cyl + wt)
```

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	34.843	2.677	13.014	5.03e-08 ***
cyl6	-3.324	1.806	-1.841	0.0927 .
cyl8	-4.532	2.781	-1.630	0.1314
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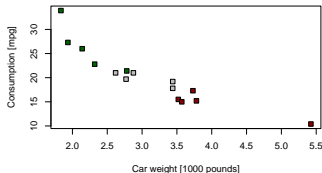
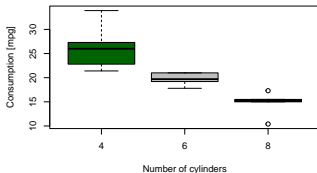
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lm(formula = mpg ~ cyl + I(wt - 3), data = mtcars2)
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	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	23.177	1.412	16.419	4.39e-09 ***
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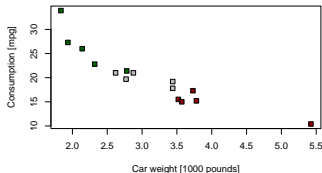
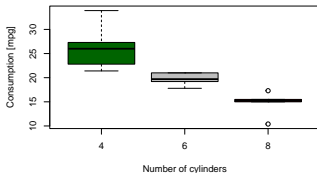
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(Intercept)	32.2238	3.5719	9.021	2.05e-06 ***
cyl1	2.6189	1.4001	1.870	0.0882 .
cyl2	-0.7053	0.9119	-0.773	0.4556
wt	-3.8886	1.1085	-3.508	0.0049 **

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(Intercept)	20.5580	0.6628	31.018	4.64e-12 ***
cyl1	2.6189	1.4001	1.870	0.0882 .
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Correlation among explanatory variables

❑ Ideal scenario

- ❑ balanced data
- ❑ uncorrelated predictors
- ❑ each coefficient β_j can be estimated separately
- ❑ interpretation of the estimated coefficients is relatively stable

❑ Typical real-data situations

- ❑ unbalanced data
- ❑ correlated predictor variables (multicollinearity)
- ❑ variance of the estimated parameters typically increases
- ❑ the interpretation of the estimated coefficients becomes rather vague

↔ in general, the value of the parameter β_j stands for a difference in the expected (conditional) value of Y which comes with a unit difference within the X_j covariate while all other predictors are kept fixed.

In practice, however, predictor variables typically change simultaneously and also, moreover, they even affect one another...

Example: Body fat vs. weight and height

□ Body fat vs. person's height

(Model 1)

```
lm(formula = fat ~ height, data = Policie)
```

Coefficients:

	Estimate	Std.Error	t value	Pr(> t)	
(Intercept)	-47.6791	23.9707	-1.989	0.0524	.
height	0.3405	0.1343	2.535	0.0146	*

□ Body fat vs. person's weight

(Model 2)

```
lm(formula = fat ~ weight, data = Policie)
```

Coefficients:

	Estimate	Std.Error	t value	Pr(> t)	
(Intercept)	-20.75217	3.42327	-6.062	2.02e-07	***
weight	0.42674	0.04266	10.003	2.51e-13	***

What about a multiple model?

□ Body fat vs. person's height and weight

(Model 3)

```
lm(formula = fat ~ height + weight, data = Policie)
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Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	16.55309	15.24621	1.086	0.2831
height	-0.24362	0.09728	-2.504	0.0158 *
weight	0.50418	0.05095	9.896	4.49e-13 ***

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- What is the estimated effect of the weight on the overall body fat?
- How well the conclusions correspond among different models?

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- How well the conclusions correspond among different models?
- The estimated correlation between the weight and height is 0.6068!

How to overcome the problems?

□ Body fat vs. person's height and weight with interactions (Model 4)

```
lm(formula = fat ~ height + weight + height:weight)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-48.604790	87.698149	-0.554	0.582
height	0.123659	0.496447	0.249	0.804
weight	1.324727	1.088637	1.217	0.230
height:weight	-0.004608	0.006106	-0.755	0.454

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- ❑ **What are the main pros/cons of the model with interactions?**

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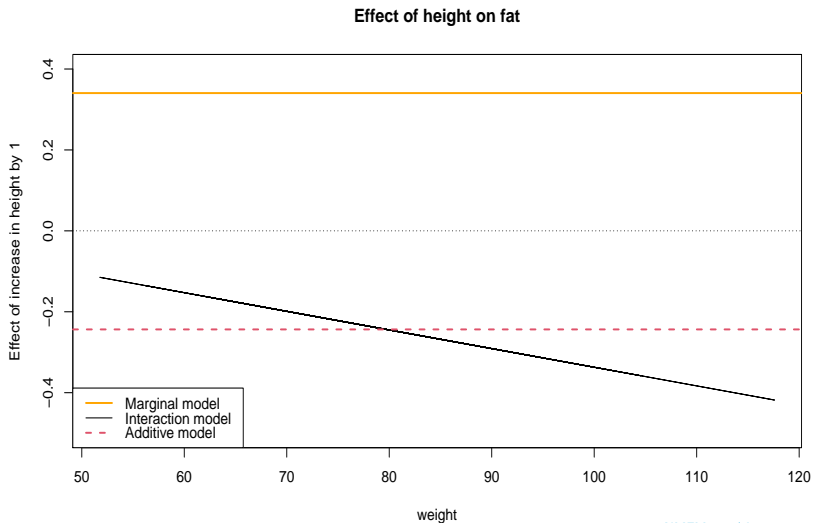
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Note, that [Model 1](#) and [Model 2](#) are principally different—they are using different set of covariates to explain the conditional expectation of `fat`. On the other hand, [Model 3](#) and [Model 4](#) are also different but both are using the same information provided by the same set of covariates. [Model 3](#) is actually a **submodel** of [Model 4](#). Moreover, both [Model 1](#) and [Model 2](#) are also submodels of [Model 3](#) or [Model 4](#).

Illustration of the Models 1, 3, and 4



Regression model with interactions: Formally

❑ Implementation in the R software

- ❑ using the expression `height:weight` (interactions only)
- ❑ using the expression `height * weight` (hierarchically well structured)
- ❑ defining new covariate as a product of `height` and `weight`

❑ Formulation within a linear regression model

- ❑ using a regression model expression: $Y \approx \beta_0 + \beta_1 X_h + \beta_2 X_w + \beta_3 X_h X_w$
- ❑ using a new covariate $Y \approx \beta_0 + \beta_1 X_h + \beta_2 X_w + \beta_3 Z$ where $Z = X_h \times X_w$

❑ More general formulations and models

- ❑ effect of height: $Y \approx \beta_0 + (\beta_1 + \beta_3 X_w) X_h + \beta_2 X_w$
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❑ Thus

- ❑ parameter β_3 can be seen as a linear function of X_w (or X_h respectively)
- ❑ more generally, β_3 is a function of X_w (or X_h respectively)
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What if $\beta_3(x) = g(x)$ for some general (smooth/continuous) function g ?

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What if $\beta_3(x) = g(x)$ for some general (smooth/continuous) function g ?

When we investigate the effect of the height on the overall amount of fat the other covariate (weight) acts as an **effect modifier** in the underlying model (and vice versa)

When to use a model with interactions?

❑ Effect modifier

If there is a belief that the effect of one specific covariate $X \in \mathbb{R}$ will be different in different sub-populations that we want/need to control for in the model by using the remaining covariates

❑ Colinearity issues

If the model design is not optimal and it is possible that some covariates may be substantially correlated (i.e., linearly dependent – multicollinearity) then the interaction(s) may help to improve the model

❑ Model interpretability

Interactions can be also used for the purpose of some better model interpretation (despite the fact that typically the interactions make the model interpretability more complex/challenging)
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Interactions are not necessarily between just two explanatory covariates (so-called **double interactions**, or **first-order interactions**). In practice, **higher-order** interactions between three and even more covariates can be implemented in the model (but the model interpretation becomes very challenging, almost impossible...)

Interpretation of the interaction term

- Consider a simple regression model with one interaction

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 (X_1 \times X_2) + \varepsilon$$

- We are primarily interested in the effect of X_1 on $E[Y|X_1, X_2]$ thus, we can rewrite the model in the equivalent form

$$Y = \beta_0 + (\beta_1 + \beta_3 X_2) X_1 + \beta_2 X_2 + \varepsilon$$

- To describe the effect of X_1 on $E[Y|X_1, X_2]$ we need to estimate/predict $(\beta_1 + \beta_3 X_2)$ which, however, depends on a (random) value of X_2 – taking (hypotetically) infinitely many values Which ones to use?
- For a given $X_2 = x_2$, the effect of X_1 on $E[Y|X_1, X_2]$ reduces to $\beta_1 + \beta_3 x_2$ which can be easily estimated ..
- For $X_2 \in \{0, 1\}$ (binary), the effect of X_1 on $E[Y|X_1, X_2]$ reduces to two distinct effects in two separate groups defined by X_2

Towards model building process...

□ Nonlinear transformations

Many different transformation functions $t \in \mathcal{G}$ can be considered within the regression model

$$Y = \beta_0 + \beta_1 t_1(X_1) + \beta_2 t_2(X_2) + \varepsilon$$

to improve its flexibility but different transformations (different choices of $t_1, t_2 \in \mathcal{G}$) change the overall model (its properties, interpretation, etc.) and the models are not directly comparable among each other

□ Linear transformations

There is a specific class of transformations that preserve most of the model qualities (if the model is well-formulated) and these are linear transformations of the form $t(x) = a + bx$, i.e.,

$$Y = \beta_0 + \beta_1(a_1 + b_1 X_1) + \beta_2(a_2 + b_2 X_2) + \varepsilon$$

for $a_1, a_2, b_1, b_2 \in \mathbb{R}$ – models under such transformations are equivalent (if $b_1 \neq 0 \neq b_2$) and can be directly compared among each other...

Linear transformations of the covariates

Typically they are used to

- ❑ to improve the stability of the estimated parameters
(e.g., measuring the distance between Prague and Brno in millimeters/kilometers)
- ❑ for better (visual) representation of the model outputs
(mostly using different units, scales, proportions for better visualization)
- ❑ to improve the interpretation of the final model
(typically, we want to have a reasonable interpretation of the intercept and interactions)

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However, it only works with a **hierarchically well-formulated** model...

- ❑ What is a **hierarchically well-formulated/structured** model?
- ❑ What are the consequences of a non-hierarchical model?

Model hierarchy

A **hierarchically well-formulated linear regression model** (often called a hierarchical or marginality-respecting model) follows a simple but important principle: If a higher-order term is included (e.g., interactions or polynomials), all corresponding lower-order terms must also be included in the model.

❑ Advantages

- ❑ linear transformations of the covariates does not effect the model
- ❑ different models are better compared within their hierarchical structure
- ❑ systematic model building procedures are well defined and work well

❑ Disadvantages

- ❑ some models can not be fitted under the restriction of hierarchy
- ❑ models with various irregularities (discontinuous, non-smooth
- ❑ sometimes it is necessary to use a model without the intercept

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↔ when fitting a linear regression model, we always need to be aware of its structure – whether we are building a model that is hierarchically well formulated or not... and depending on the model we have different tools available for the fitting process and the consecutive inference as well

Summary

❑ Models with interactions

- ❑ they may help to overcome some issues with the covariates
- ❑ they may improve the overall flexibility of the model
- ❑ interpretation of the model usually becomes more challenging

❑ Linear transformations of the covariates

- ❑ they help with the model/computational stability
- ❑ when used wisely, they improve the interpretation of the final model
- ❑ they require a hierarchically well formulated model to work properly

❑ Hierarchically well-formulated (HWF) models

- ❑ have some specific advantages and disadvantages
- ❑ inference in a hierarchical model is more straightforward
- ❑ however, some practical applications may require a non-hierarchical model

❑ Key message

- ❑ some models are in a mutual relationship (i.e., model vs. submodel)
- ❑ a sequence of submodels is the main tool for the model building purposes
- ❑ for a proper comparisons of the submodels, we need HWF models